

# *The Determination of a Trend in a Multi-Scale Problem*

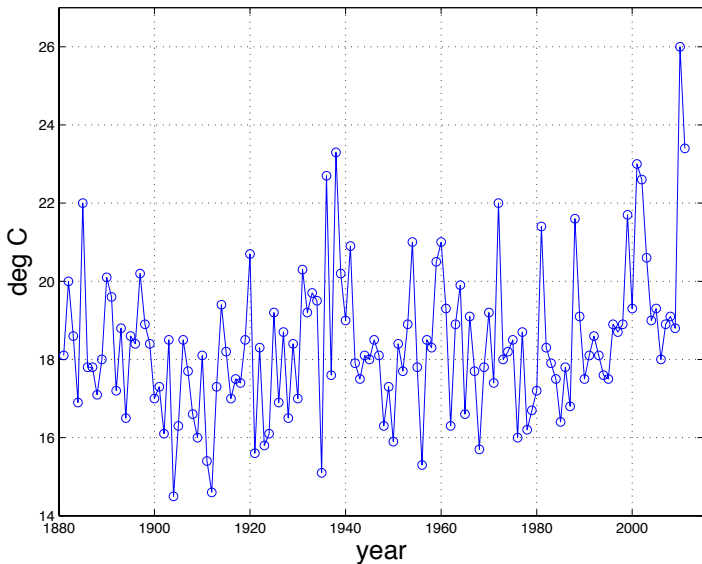
## *How Warm is it Getting?*

JUAN M. RESTREPO

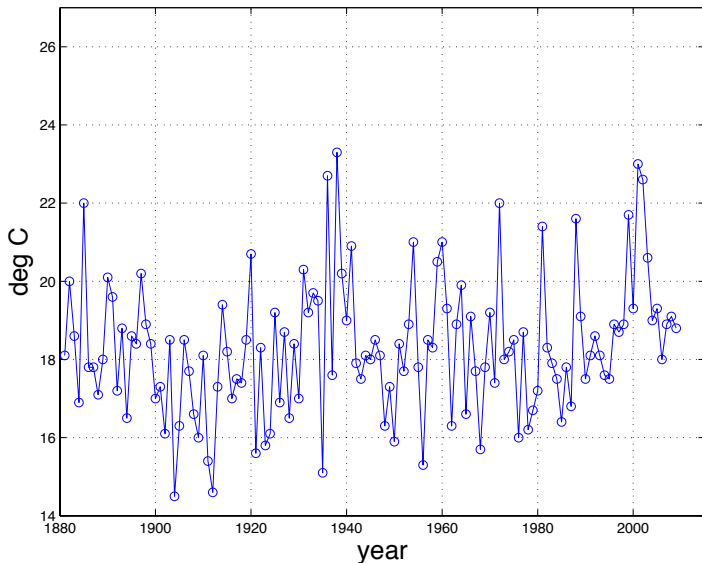
*Department of Mathematics, Department of Statistics  
and Physical Oceanography Oregon State University*

May 21, 2018

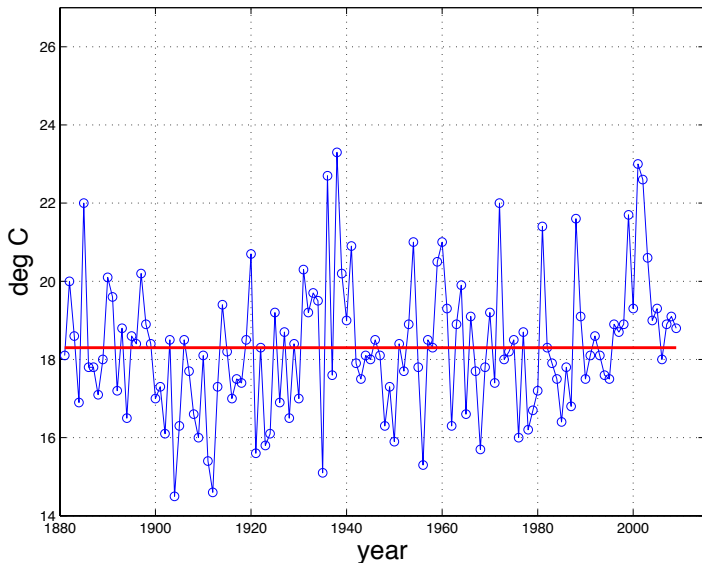
# Moscow's Summer Temperatures, 1881-2011



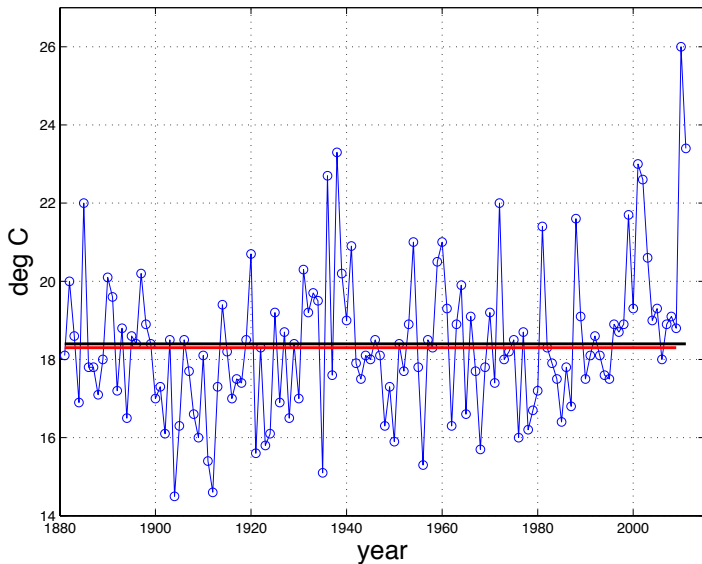
# Moscow's Summer Temperatures, 1881-2009



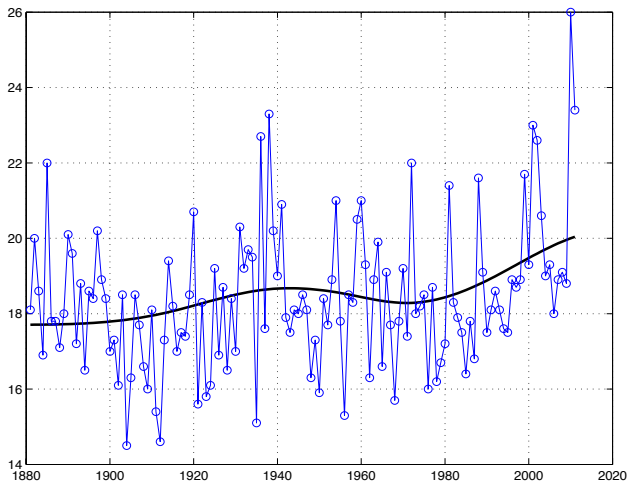
# Moscow's Summer Temperatures, 1881-2009



# Moscow's Summer Temperatures, 1881-2011



# Something Must Account for Changing Mean



*Increase of Extreme Events in a Warming World (PNAS 44, 2011), by Rahmstorf and Coumou.*

## The Trend Problem:

Define a set of simple universal rules with which to compute an underlying *tendency*, given a finite (non-stationary/multi-scale) data set.

Joint work with

**Shankar Venkataramani (U. Arizona)**

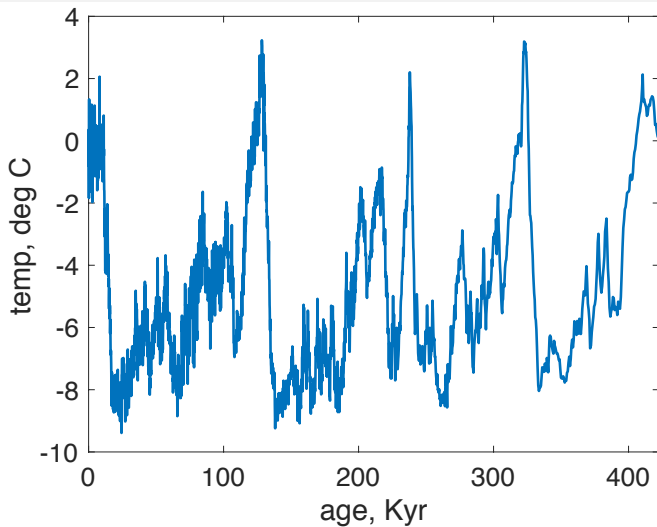
**H. Flaschka (U. Arizona)** and

**D. Comeau (U. Arizona)**



National Science Foundation  
WHERE DISCOVERIES BEGIN

# A Climate Signal...



Vostok Ice Core data, Temperature



Given a finite-time time series  $Y(i)$ ,  $i = 1, 2, \dots, N$ ,

The **Tendency**  $T(i)$  is an *Executive Summary* of  $Y(i)$

- Captures essentials of histogram in the abscissa of  $Y(i)$ ; *and*
- Most essential multi scale information, derived from ordinate of  $Y(i)$ .

The **Empirical Uncertainty**  $U(i) := Y(i) - T(i)$

- is simple entropically,
- The histogram of  $U(i)$ , is easy to parametrize

Given a finite-time time series  $Y(i)$ ,  $i = 1, 2, \dots, N$ ,

The **Tendency**  $T(i)$  is an *Executive Summary* of  $Y(i)$

- Captures essentials of histogram in the abscissa of  $Y(i)$ ; *and*
- Most essential multi scale information, derived from ordinate of  $Y(i)$ .

The **Empirical Uncertainty**  $U(i) := Y(i) - T(i)$

- is simple entropically,
- The histogram of  $U(i)$ , is easy to parametrize

## General Procedure:

- Find a decomposition  $Y(i) = B^D + \sum_{j=1}^D R^j(i)$
- Apply tendency criteria to pick  $T(i) := B^D + \{R^j(i)\}_S, i = 1, \dots, N.$ 
  - $B^D$  is a constant,
  - $\{R^j(i)\}_S$  is a function made up of a combination of  $S$  *rotations*.

The choice of decomposition is motivated by the

- Be non-parametric.
- Ability to handle **multi-scale** nature of a signal.
- Be lossless.

# The Intrinsic Time Decomposition (ITD)

Given a sequence of real numbers  $\{Y(i)\}_{i=1}^N$ ,

$$Y(i) = B^D + \sum_{j=1}^D R^j(i)$$

where

$$B^j(i) = B^{j+1}(i) + R^{j+1}(i), \quad j = 0, \dots, D,$$

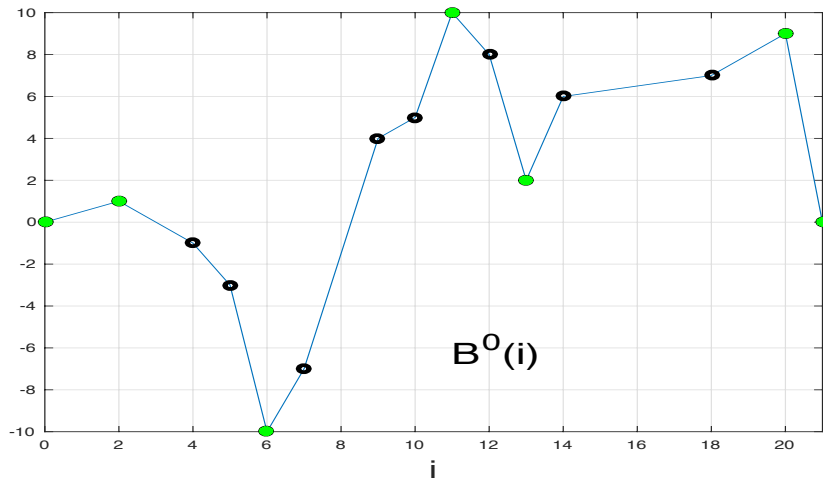
and

$$B^0(i) : = Y(i).$$

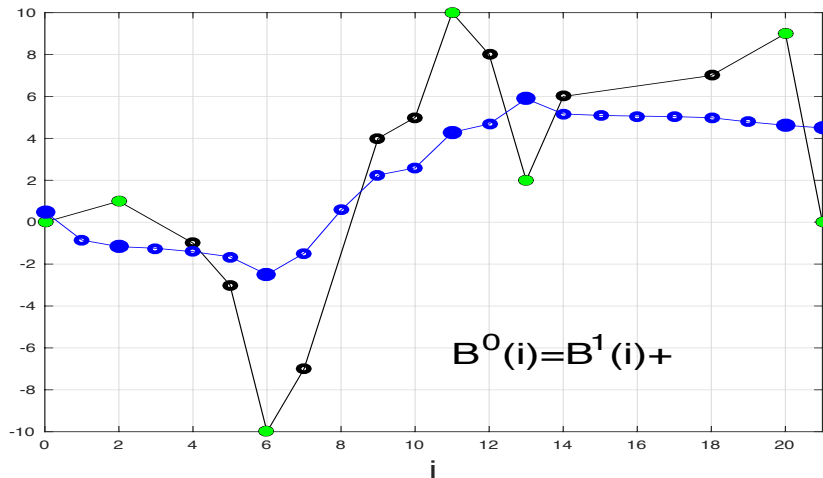
$B^j$  are called *BASELINES*, and  $R^j$  are called *ROTATIONS*.

Frei and Osorio, Proc. Roy. Soc. London, (2006).

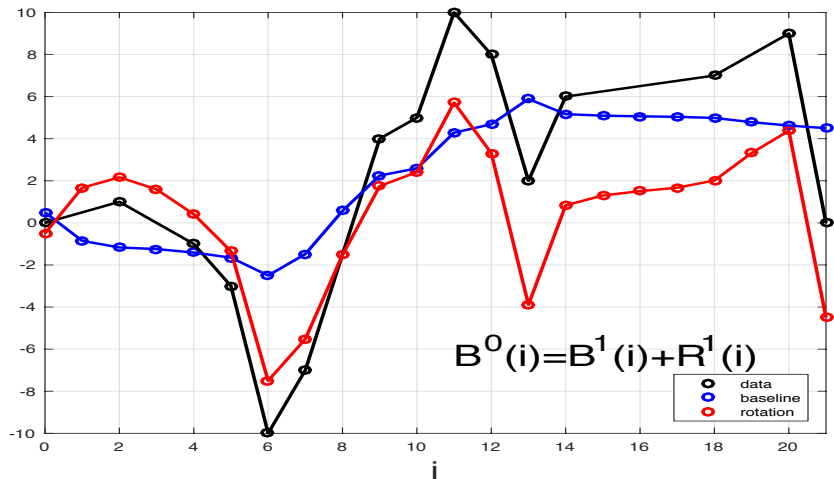
# The Intrinsic Time Decomposition (ITD)



# The Intrinsic Time Decomposition (ITD)

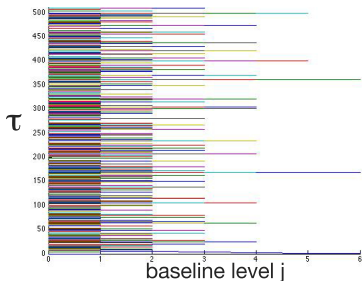
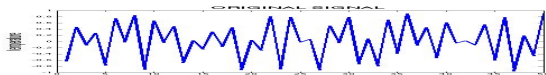


# The Intrinsic Time Decomposition (ITD)



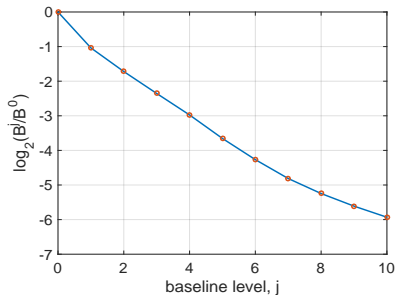
# ITD Decomposition: No Low Frequency Slowdown

All Extrema Random Signal  $Y = (-1)^i |z_i|$ ,  $z_i$  from  $\mathcal{N}(\sigma = 4)$



Extrema spacing  $\tau^0 / \tau^j$ ;

$$\tau^j / \tau^{j+1} \approx -0.41$$



$\|B^j\| / \|Y\|$

$$\|B^j\| / \|B^{j+1}\| \approx -0.61.$$



## How does the ITD (and EMD) work?

$$\mathcal{E}[B^j] := \{S^j, b^j\}.$$

$\{S^j\}_1^{n_j}$  be locations of extrema of **baselines**, with values  $b^j$ .

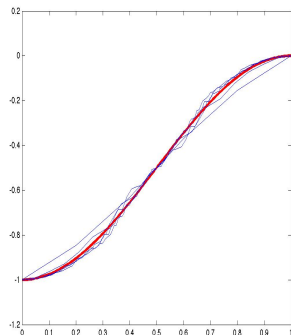
ITD:

$$\{S^{j+1}, b^{j+1}\} = \mathcal{E}[(\mathbb{I} + M^j)b^j].$$

# Self Similar Spectrum

ITD:  $\{s^{j+1}, b^{j+1}\} = \mathcal{E}[(\mathbb{I} + M^j)b^j]$ .

- Spectrum( $\mathbb{I} + M^j$ ):  $\lambda_k^j = \cos^2(\pi k/n)$ ,
- e' value 1 corresponding to right e' vector consisting of all 1,
- e' value 0 corresponding to right e' vector consisting of  $x_k = (-1)^k$ .



**Spectrum( $\mathbb{I} + M^j$ ),  
for all levels  $j$**

# Universality and Decay of the $\ell_2$ of the Baselines

$$\{S^{j+1}, b^{j+1}\} = \mathcal{E}[(I + M^j)b^j]$$

at  $j + 1$ , with  $S^{j+1} \subseteq S^j$ .

$$v := M^j b \sim \frac{1}{4}B''(x) + \frac{1}{2}p^j(x)B'(x).$$

$$v_k = \frac{1}{4}(b_{k-1} - 2b_k + b_{k+1}) + \frac{p_k^j}{4}(b_{k+1} - b_{k-1}),$$

$p_k^j \in (-1, 1)$  given by

$$p_k^j = \frac{2\tau_k^j - \tau_{k-1}^j - \tau_{k+1}^j}{\tau_{k+1}^j - \tau_{k-1}^j}.$$

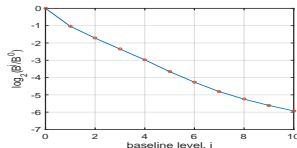
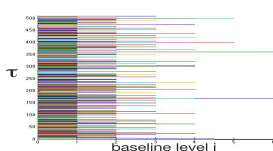
Hence,  $(I + M^j)b^j$  is FTCD approximation of

$$\frac{\partial}{\partial t} B = \frac{1}{4} \frac{\partial^2}{\partial x^2} B + \frac{1}{2} p^j(x) \frac{\partial}{\partial x} B = \frac{1}{4w^j(x)} \frac{\partial}{\partial x} \left[ w^j(x) \frac{\partial B}{\partial x} \right],$$

$$w^j(x) = \exp \left[ 2 \int_0^x p^j(t) dt \right].$$

# Distancing of Extremas and Decay of Baselines, with Level $j$

All Extrema Random Signal  $Y = (-1)^i |z_i|$ ,  $z_i$  from  $\mathcal{N}(\sigma = 4)$



Extrema spacing  $\tau^0 / \tau^j$ ;

$$\|B^j\| / \|Y\|$$

Model:  $b_k^j \approx \mu^j (-1)^k + \alpha^j n_k$ .

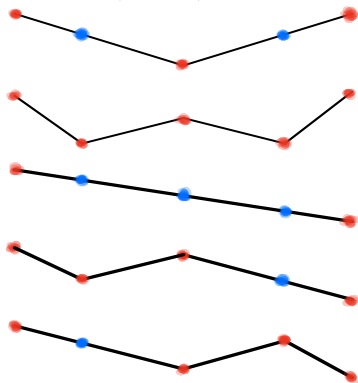
Since  $(I + M^j)(-1)^k = 0$  then  $b^{j+1} = \alpha^j \mathcal{E}((I + M^j)\mathbf{n})$ , where  $\mathbf{n} = n_k$  is a vector of independent normal variates.

# Estimating $\mu$ and $\alpha$

Model:  $b_k^j \approx \mu^j (-1)^k + \alpha^j n_k$ .

Let  $\mathbf{x} = (x_1, x_2, x_3)$ , consecutive entries of the vector  $(I + M^j)\mathbf{n}$ , then

*count ways  $x_i$  can be extrema* ●



Let  $x_1, x_2$  and  $x_3$  consecutive entries of the vector  $(I + M^j)\mathbf{n}$ , then

$$\mathbf{x} = \begin{pmatrix} x_1 \\ x_2 \\ x_3 \end{pmatrix} \approx \frac{1}{4} \begin{pmatrix} 1-p_1 & 2 & 1+p_1 & 0 & 0 \\ 0 & 1-p_2 & 2 & 1+p_2 & 0 \\ 0 & 0 & 1-p_3 & 2 & 1+p_3 \end{pmatrix} \begin{pmatrix} n_1 \\ n_2 \\ n_3 \\ n_4 \\ n_5 \end{pmatrix} \equiv \mathbf{A}\mathbf{n}.$$

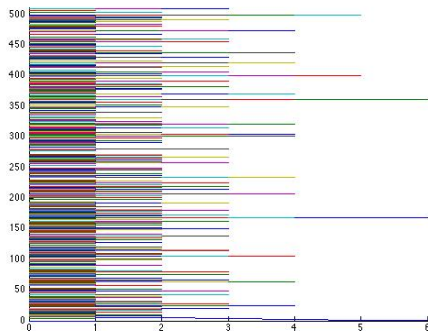
where  $n_i$  independent normal,  $p_i$  are realizations. So  $x_1, x_2$  and  $x_3$  are jointly Gaussian with mean zero and covariance  $\Sigma(p_1, p_2, p_3) = \mathbf{A}\mathbf{A}^T$  and the (conditional) joint density of  $x_1, x_2$  and  $x_3$  is given by

$$p(x_1, x_2, x_3 | p_1, p_2, p_3) = \frac{1}{\sqrt{8\pi^3 \text{Det}(\Sigma(p_1, p_2, p_3))}} \exp \left[ -\frac{1}{2} \mathbf{x}^T \Sigma(p_1, p_2, p_3)^{-1} \mathbf{x} \right].$$

# The Joint Density of the Spacing

Let  $l_k = \tau_{k+1} - \tau_k$ , then  $p_k = \frac{l_k - l_{k-1}}{l_k + l_{k-1}}$ . Assume every site has a constant probability of being an extremum, and no correlation between neighbors.

Guess  $\hat{p}$  is exponentially distributed (mean = 1)



The joint density of  $p_1, p_2$  and  $p_3$ :

$$\rho(p_1, p_2, p_3) = \frac{128(1-p_1)(1+p_2)(1-p_2)(1+p_3)}{(3-p_1+p_2+p_1p_2)^3(3-p_2+p_3+p_2p_3)^3}.$$

and

$$\begin{aligned} \Sigma_{av} &= \int_{-1}^1 dp_2 \int_{-1}^1 dp_1 \int_{-1}^1 dp_3 \Sigma(p_1, p_2, p_3) \rho(p_1, p_2, p_3) \\ &= \begin{pmatrix} 0.41666\dots & 0.25 & 0.058227\dots \\ 0.25 & 0.41666\dots & 0.25 \\ 0.058227\dots & 0.25 & 0.41666\dots \end{pmatrix}. \end{aligned}$$

We now obtain the joint density of  $x_1, x_2$  and  $x_3$  to get

$$p_{av}(x_1, x_2, x_3) \approx \frac{1}{\sqrt{8\pi^3 \text{Det}(\Sigma_{av})}} \exp \left[ -\frac{1}{2} \mathbf{x}^T \Sigma_{av}^{-1} \mathbf{x} \right].$$



We can also compute the mean and the variance of the distribution of the minima (**maxima**).

$$\mu = \frac{1}{\beta} \int_{-\infty}^{\infty} dx_3 \int_{x_3}^{\infty} dx_2 \int_{-\infty}^{x_2} dx_1 x_2 p(x_1, x_2, x_3) = 0.483883 \dots$$

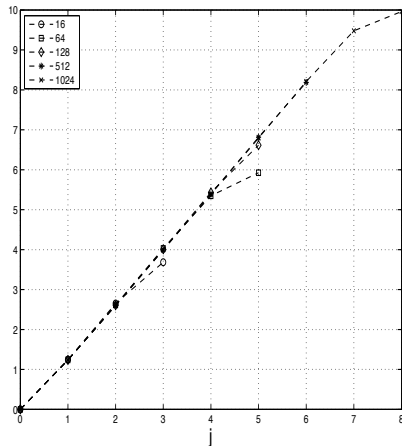
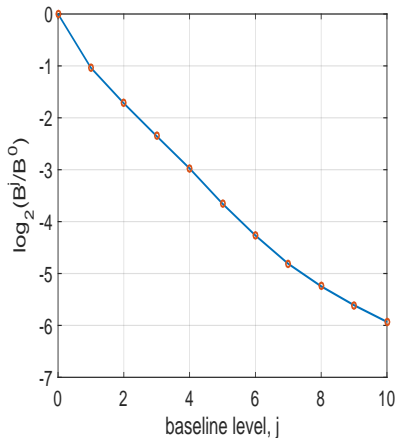
and

$$\alpha^2 = \frac{1}{\beta} \int_{-\infty}^{\infty} dx_3 \int_{x_3}^{\infty} dx_2 \int_{-\infty}^{x_2} dx_1 x_2^2 p(x_1, x_2, x_3) - \mu^2 = 0.302712 \dots$$

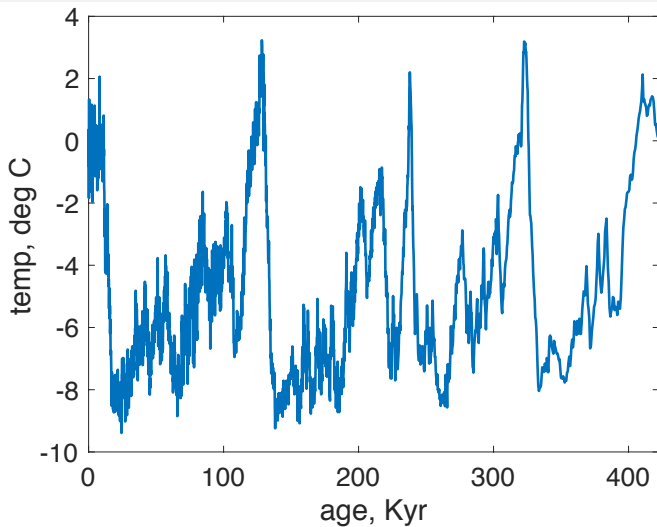
From this, we obtain

$$(I + M^j)\mathbf{n} \approx \mu(-1)^k + \alpha\mathbf{n}' = 0.483883 \times (-1)^k + 0.550193\mathbf{n}'.$$

|             | $B^j/B^0$ DECAY RATE | $\tau^0/\tau^j$ DECAY RATE |
|-------------|----------------------|----------------------------|
| THEORETICAL | -0.55                | -0.48                      |
| NUMERICAL   | -0.61                | -0.41                      |

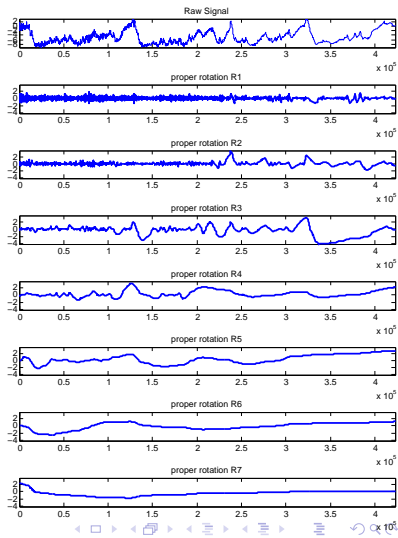
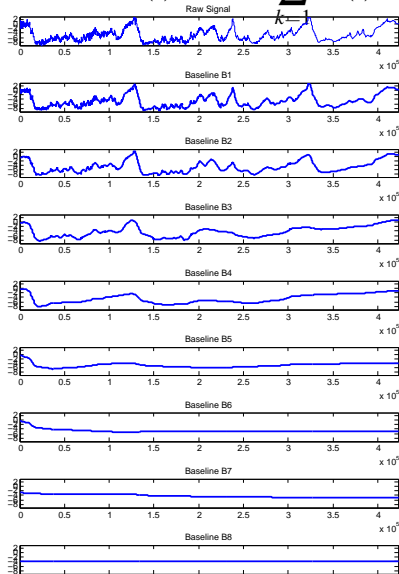


# Example Calculation



Vostok Ice Core data, Temperature

$$Y(i) = B^D + \sum_{k=1}^D R^k(i), \quad B^{j+1}(t) + R^{j+1}(i) = B^j(i).$$



# Finding the Tendency

- Find ITD:

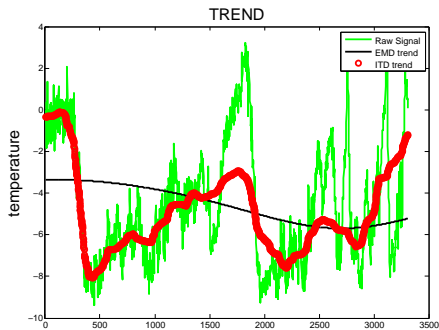
$$Y(i) = B^D + \sum_{j=1}^D R^j(i),$$

$$B^j(i) = B^{j+1}(i) + R^{j+1}(i)$$

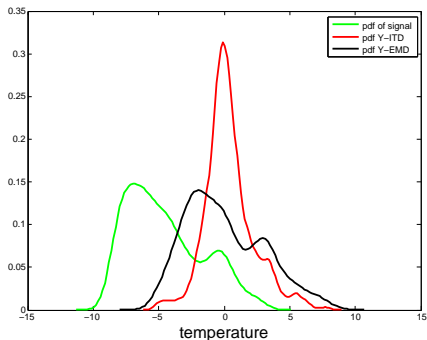
- Find Tendency (picking  $j^*$  baseline)

$$T(i) := B^{j^*}(i)$$

# The Tendency $T(i)$ , the EMD, and the Vostok signal $Y(i)$



Time Series



The Histograms

# Find Tendency

Choosing  $j^*$  among the baselines  $\{B^j(i)\}_{j=1}^D$ :

$$T(i) := B^{j^*}(i)$$

## The ABSISSA information:

- For  $j = 1, \dots, D$  compute  $F^j := \text{histogram}[Y(i) - B^j(i)]$
- Determine the **Symmetry**  $s^j$  of  $F^j$  via percentiles:

$$s^j := \frac{Pr_{75}^j - 2Pr_{50}^j - Pr_{25}^j}{(Pr_{75}^j - Pr_{25}^j)}$$

Choosing  $j^*$  among the **baseline**  $\{B^j(i)\}_{j=1}^D$ :

$$T(i) := B^{j^*}(i)$$

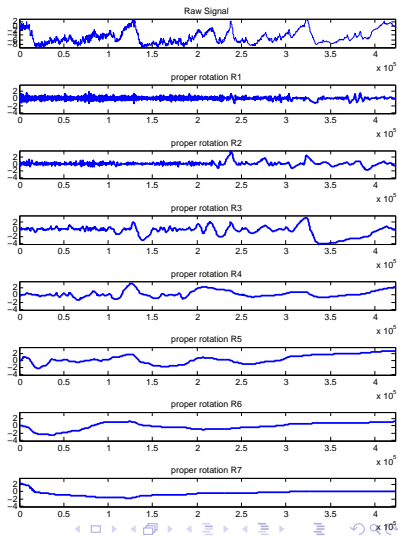
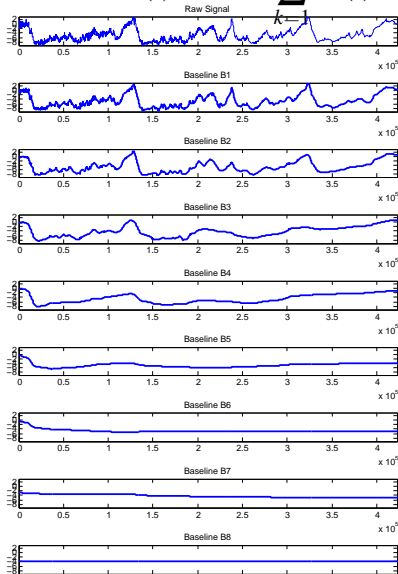
**The ORDINATE information:**

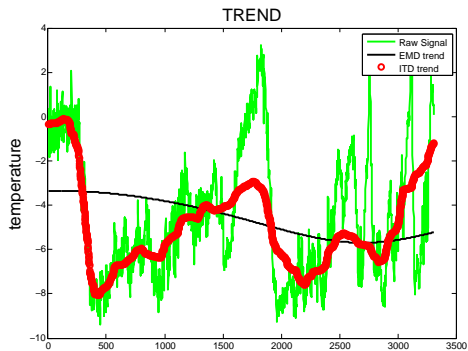
- Compute the **Complexity**  $c_j$  vector

$$c^j := \text{corr}(B^j, R^j)$$

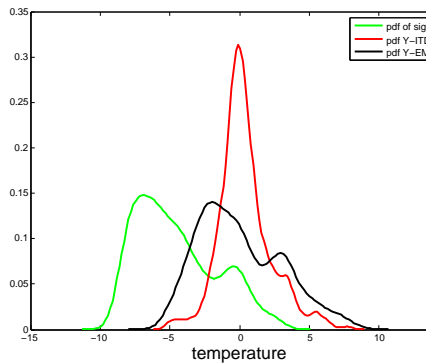


$$Y(i) = B^D + \sum_{k=1}^D R^k(i), \quad B^{j+1}(t) + R^{j+1}(i) = B^j(i).$$





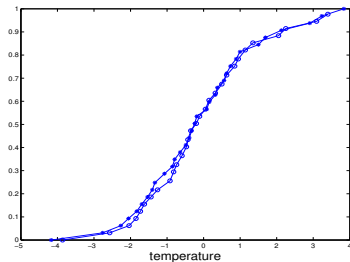
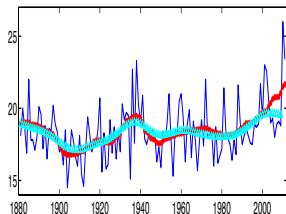
Time Series



The Histograms

# Analysis of the Moscow Data

*Our analysis confirms Rahmstorf and Coumou's guess: the mean temperature increased, but not its variance:*



# Further Information

Juan M. Restrepo

<http://www.math.oregonstate.edu/~restrepo>

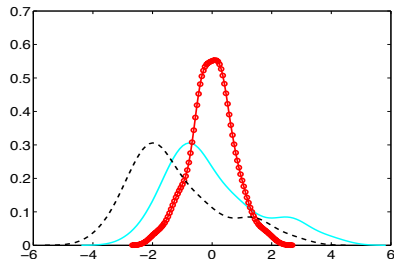
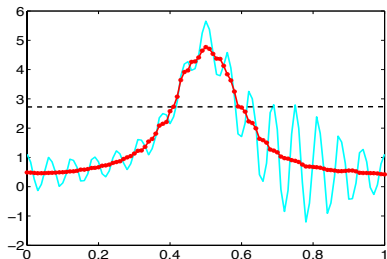


National Science Foundation  
WHERE DISCOVERIES BEGIN

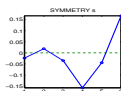
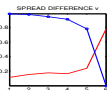
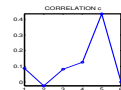
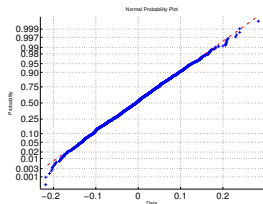
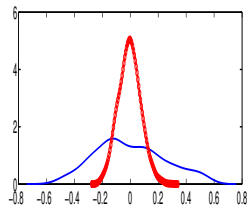
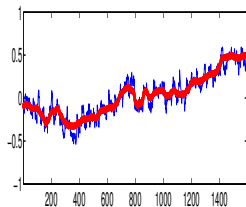
# A Multiscale Signal

$$Y(i) = \frac{1}{1.5 + \sin(2\pi t)} \cos[32\pi t_i + 0.2 \cos(64\pi t_i)] + \frac{1}{(1.2 + \cos(2\pi t_i))};$$

for  $t_i \in [0 : 0.0025 : 1]$ .



# Ocean Temperatures



# The Composite Case

